

# North Atlantic climate far more predictable than models imply

Article

Accepted Version

Smith, D. M., Scaife, A. A., Eade, R., Athanasiadis, P., Bellucci, A., Bethke, I., Bilbao, R., Borchert, L. F., Caron, L.-P., Counillon, F., Danabasoglu, G., Delworth, T., Doblas-Reyes, F. J., Dunstone, N. J., Estella-Perez, V., Flavoni, S., Hermanson, L., Keenlyside, N., Kharin, V., Kimoto, M., Merryfield, W. J., Mignot, J., Mochizuki, T., Modali, K., Monerie, P.-A. ORCID: https://orcid.org/0000-0002-5304-9559, Müller, W. A., Nicoli, D., Ortega, P., Pankatz, K., Pohlmann, H., Robson, J. ORCID: https://orcid.org/0000-0002-3467-018X, Ruggieri, P., Sospedra-Alfonso, R., Swingedouw, D., Wang, Y., Wild, S., Yeager, S., Yang, X. and Zhang, L. (2020) North Atlantic climate far more predictable than models imply. Nature, 583. 7818. ISSN 0028-0836 doi: https://doi.org/10.1038/s41586-020-2525-0 Available at https://centaur.reading.ac.uk/91527/

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To link to this article DOI: http://dx.doi.org/10.1038/s41586-020-2525-0

Publisher: Nature Publishing Group



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## Climate models underpredict North Atlantic atmospheric 2 circulation changes

- <sup>3</sup> D. M. Smith<sup>1</sup>, A. A. Scaife<sup>1,2</sup>, R. Eade<sup>1</sup>, P. Athanasiadis<sup>3</sup>, A. Bellucci<sup>3</sup>, I. Bethke<sup>4,5</sup>, R. Bilbao<sup>6</sup>, L.
- <sup>4</sup> F. Borchert<sup>7</sup>, L.-P. Caron<sup>6</sup>, F. Counillon<sup>4,5</sup>, G. Danabasoglu<sup>8</sup>, T. Delworth<sup>9</sup>, F. J. Doblas-Reyes<sup>6,10</sup>,
- <sup>5</sup> N. J. Dunstone<sup>1</sup>, V. Estella-Perez<sup>7</sup>, S. Flavoni<sup>7</sup>, L. Hermanson<sup>1</sup>, N. Keenlyside<sup>4,5</sup>, V. Kharin<sup>11</sup>, M.
- <sup>6</sup> Kimoto<sup>12</sup>, W. J. Merryfield<sup>11</sup>, J. Mignot<sup>7</sup>, T. Mochizuki<sup>13,14</sup>, K. Modali<sup>15</sup>, P.-A. Monerie<sup>16</sup>, W. A.
- <sup>7</sup> Müller<sup>15</sup>, D. Nicolí<sup>3</sup>, P. Ortega<sup>6</sup>, K. Pankatz<sup>17</sup>, H. Pohlmann<sup>15,17</sup>, J. Robson<sup>16</sup>, P. Ruggieri<sup>3</sup>, R.
- <sup>8</sup> Sospedra-Alfonso<sup>11</sup>, D. Swingedouw<sup>18</sup>, Y. Wang<sup>4</sup>, S. Wild<sup>6</sup>, S. Yeager<sup>8</sup>, X. Yang<sup>9</sup> and L. Zhang<sup>9</sup>
- <sup>9</sup> <sup>1</sup>Met Office Hadley Centre, FitzRoy Road, Exeter EX1 3PB, UK
- <sup>10</sup> <sup>2</sup>College of Engineering, Mathematics and Physical Sciences, Exeter University, UK
- <sup>11</sup> <sup>3</sup>Centro Euro-Mediterraneo sui Cambiamenti Climatici, Bologna, Italy
- <sup>12</sup> <sup>4</sup>Nansen Environmental and Remote Sensing Center and Bjerknes Centre for Climate Research,
- 13 Bergen, Norway
- <sup>14</sup> <sup>5</sup>Geophysical Institute, University of Bergen and Bjerknes Centre for Climate Research, Bergen,
- 15 Norway
- <sup>16</sup> <sup>6</sup>Barcelona Supercomputing Center, Jordi Girona 29 08034 Barcelona, Spain
- <sup>17</sup> <sup>7</sup>Sorbonne Universités, LOCEAN Laboratory, Institut Pierre Simon Laplace (IPSL), Paris, France
- <sup>18</sup> <sup>8</sup>National Center for Atmospheric Research, Boulder, CO, USA
- <sup>19</sup> <sup>9</sup>Geophysical Fluid Dynamics Laboratory, Princeton University, Princeton, NJ, USA
- <sup>20</sup> <sup>10</sup>*ICREA*, *Barcelona*, *Spain*
- <sup>21</sup> <sup>11</sup>Canadian Centre for Climate Modelling and Analysis, Environment and Climate Change Canada,

- 22 Victoria, British Columbia, Canada
- <sup>23</sup> <sup>12</sup> Atmosphere and Ocean Research Institute, University of Tokyo, Kashiwa, Japan
- <sup>24</sup> <sup>13</sup>Department of Earth and Planetary Sciences, Kyushu University, Fukuoka, Japan
- <sup>25</sup> <sup>14</sup>Japan Agency for Marine-Earth Science and Technology, Yokohama, Japan
- <sup>26</sup> <sup>15</sup> Max-Planck-Institut für Meteorologie, Bundesstraße 53, 20146 Hamburg, Germany
- <sup>27</sup> <sup>16</sup>National Centre for Atmospheric Science, Department of Meteorology, University of Reading,
- <sup>28</sup> Reading RG6 6BB, UK
- <sup>29</sup> <sup>17</sup>Deutscher Wetterdienst, Bernhard-Nocht-Str. 76, Hamburg, Germany
- <sup>30</sup> <sup>18</sup>CNRS-EPOC, Université de Bordeaux, Pessac, France

<sup>31</sup> Corresponding author: Doug Smith, doug.smith@metoffice.gov.uk

#### 32 Abstract

Quantifying signals and uncertainties in climate models is essential for climate change de-33 tection, attribution, prediction and projection<sup>1-3</sup>. Although inter-model agreement is high 34 for large-scale temperature signals, dynamical changes in atmospheric circulation are very 35 uncertain<sup>4</sup>, leading to low confidence in regional projections especially for precipitation over 36 the coming decades<sup>5,6</sup>. Furthermore, model simulations with tiny differences in initial condi-37 tions suggest that uncertainties may be largely irreducible due to the chaotic nature of the cli-38 mate system<sup>7–9</sup>. However, climate projections are difficult to verify until further observations 39 become available. Here we assess retrospective climate predictions of the last six decades 40

and show that decadal variations in north Atlantic winter atmospheric circulation are highly 41 predictable. Crucially, climate models underestimate the predictable signal by an order of 42 magnitude and skill is achieved despite a lack of agreement between individual model simu-43 lations. Consequently, skilful climate predictions of European and eastern North American 44 winters are possible but require 100 times more ensemble members than would perfect mod-45 els and post-processing to overcome underestimated teleconnections. Our results highlight 46 the pressing need to understand why the signal-to-noise ratio is too small in climate models<sup>10</sup>, 47 and the extent to which correcting this model error would reduce uncertainties in regional 48 climate change on timescales beyond a decade. 49

Global climate models are used extensively to understand the drivers of past climate variabil-50 ity and change, and to predict what is likely to happen in the future<sup>1-3</sup>. Underpinning this is a need 51 for accurate estimates of signals and associated uncertainties in climate model simulations in order 52 to distinguish between different causes of past climate change, and to provide reliable confidence 53 limits on future projections. Uncertainties are typically partitioned into three sources<sup>11</sup>: scenario 54 uncertainty arising from an imperfect knowledge of external forcing factors, including changes 55 in greenhouse gases, ozone, anthropogenic and volcanic aerosols, and solar irradiance; modelling 56 uncertainty arising from the fact that different models respond differently to the same radiative 57 forcing; and internal variability of climate that would occur in the absence of any external forcing. 58

<sup>59</sup> Climate projections for many regions are currently highly uncertain, especially for atmo-<sup>60</sup> spheric circulation<sup>4,12</sup> and related impacts, including precipitation<sup>5,6</sup>. This is particularly well

illustrated by the fact that modelling<sup>13,14</sup> and internal variability<sup>7,8</sup> uncertainties are each large 61 enough to allow opposite projections of European winters, especially for the coming decades. 62 Whilst modelling uncertainties might be reduced as models improve, internal variability uncer-63 tainties have been interpreted to be largely irreducible<sup>7–9</sup> suggesting that confident projections of 64 European winters may never be possible. However, such conclusions assume that signals and un-65 certainties diagnosed from climate models are correct. Although multi-decadal and longer climate 66 projections are difficult to verify until future observations become available, signals over the first 67 10 years can be more robustly evaluated using retrospective decadal predictions (hereafter referred 68 to as hindcasts). 69

We use a very large multi-model ensemble of decadal hindcasts from the Coupled Model 70 Intercomparison Project (CMIP) phases  $5^{15}$  and  $6^{16}$ . We focus on the boreal winter period (De-71 cember to March) averaged over forecast years 2 to 9 to avoid seasonal to annual predictability 72 and focus on decadal timescales. We use hindcasts starting each year over the period 1960 to 2005 73 from 6 CMIP5 and 8 CMIP6 modelling systems, giving a total of 169 ensemble members which 74 are weighted equally (see Methods, Table1). Hence our total hindcast dataset comprises 77,740 75 (46 start dates times 169 ensemble members times 10 years) years of model integrations to provide 76 robust statistics. 77

To compare with uncertainties in climate projections<sup>5, 7, 8, 13, 14</sup> we focus on European winters which are largely controlled by the North Atlantic Oscillation (NAO), the leading mode of atmospheric circulation variability in the north Atlantic<sup>17</sup>. The NAO represents the meridional gradient in mean sea level pressure (mslp), typically measured as the difference in pressure between the Azores and Iceland. Its positive (negative) phase reflects an increased (reduced) pressure gradient driving stronger (weaker) mid-latitude westerly winds with increased (reduced) storminess, and a northward (southward) shift of the jet stream. Impacts of the NAO are characterised by a quadrupole pattern, with a positive (negative) NAO driving warmer, wetter (colder, drier) conditions in northern Europe and south-east North America along with colder, drier (warmer, wetter) conditions in southern Europe and north-east North America.

We assess skill using two different measures (see Methods): anomaly correlation coefficient 88 (ACC) which measures the phase of variability, and mean-squared-skill-score (MSSS) which mea-89 sures the amplitude of variability. We find significant skill for decadal predictions of winter mslp in 90 most regions, including the north Atlantic, when measured by the ACC between the 169-member 91 ensemble mean and observations (Figure 1a). However, skill is much lower especially in the north 92 Atlantic when measured by the MSSS or the ACC of a smaller (10-member, typical of individual 93 prediction systems<sup>16</sup>) ensemble mean (Figure 1 b and c). Timeseries from the observations and 94 each model ensemble member consist of a predictable component (the signal) and unpredictable 95 internal variability (the noise). The discrepancy in skill between ACC and MSSS, and the need for 96 a large ensemble, arise because the signal-to-noise ratio is too small in the models compared to 97 observations<sup>10, 18, 19</sup>. Hence, skill is low in a 10-member ensemble mean because a larger ensemble 98 is required to reduce the noise and extract the predicted signal. In contrast, the signal resulting from 99 a large ensemble mean may capture the correct phase of observed variability giving a significant 100 ACC, but its amplitude will be much too small resulting in a low MSSS. 101

Errors in the signal-to-noise ratio can be quantified by comparing the predictable components (the predictable fraction of the total variability) in observations and models. The ratio of predictable components<sup>10, 18, 20</sup> (RPC, see Methods) is expected to be one for a perfect forecasting system; values greater than one show where the signal-to-noise ratio is erroneously too small in models. Consistent with differences in ACC and MSSS we find RPC is greater than one almost everywhere where there is skill in ACC, and especially in the north Atlantic (Figure 1d).

The NAO exhibits marked decadal variability<sup>21</sup> with a strong increase from the 1960s to the 108 1990s and a decrease thereafter (Figure 2a, black curve). The raw ensemble mean forecast shows 109 virtually no signal (Figure 2a, red curve), and the observations generally lie within the model 110 uncertainties (shading showing the 5-95% range diagnosed from the ensemble spread), although 111 the extreme values in the early 1960s and late 1980s are not well-captured by models in agreement 112 with other studies<sup>22,23</sup>. Taken at face value, as is done for climate projections<sup>5,7,8,14</sup>, the small 113 model signal and much larger spread would imply little ability to predict the NAO and a large 114 component of unpredictable internal variability. However, by comparing with observations we find 115 significant correlation skill of the ensemble mean (ACC=0.48, p=0.02), while persistence provides 116 a poor forecast (ACC=0.1). Hence, skilful climate model predictions of the NAO are possible using 117 the ensemble mean, but the signal-to-noise ratio is too small (RPC=4.2) and its variance must be 118 calibrated to provide realistic forecasts<sup>19</sup>. 119

Rescaling the ensemble mean time-series to have the same variance as the observations reveals that the predictions do capture the observed increase from the 1960s to 1990s and decrease

thereafter (Figure 2b). However, even with 169 ensemble members (Figure 2b thin red curve) 122 there are large interannual variations that are not expected or observed in 8-year rolling means. We 123 therefore create a larger lagged ensemble by taking the average of the four latest forecasts avail-124 able at each start date (giving 676 members, Figure 2b thick red curve, see Methods). This reveals 125 that the NAO is highly predictable on decadal timescales (ACC=0.79, p< 0.01) in stark contrast to 126 the lack of predictability implied by the standard interpretation of raw model output (Figure 2a). 127 Importantly, the signal-to-noise ratio is much too small in the models (RPC=11, p=0.02). The 128 total 8-year variability of the NAO in individual model members (standard deviation = 1.7 to 2.6 129 hPa, 5-95% range, year 2-9 hindcasts) is not significantly different to the observations (2.4hPa). 130 Hence the predictable signal (see Methods) is underestimated by an order of magnitude in the 131 model ensemble. Since the standard error of the ensemble mean is reduced by the square root of 132 the ensemble size, the ensemble required to extract the signal is 100 times larger than it would be 133 for perfect models. 134

The fact that the NAO signal is much too weak in models implies that the impacts of the 135 NAO will be underestimated relative to other factors such as greenhouse gases. Hence in regions 136 influenced by the NAO the ensemble mean will not reflect the true balance of driving factors and 137 simply inflating its variance to be the same as observed will not correct the error. A potential so-138 lution is to post-process the model output by selecting a subset of (20) ensemble members from 139 the lagged ensemble (of 676 members) whose simulated NAO is closest in sign and magnitude 140 to the ensemble mean NAO after adjusting this to take into account the underestimated signal. 141 These members contain close-to the correct magnitude of the forecast NAO whilst retaining influ-142

ences from greenhouse gases and other sources. We refer to this procedure as "NAO-matching"
(see Methods) and note that it builds on previous techniques<sup>24,25</sup> by using the models as much as
possible instead of observed relationships which may not be causal or robust.

We investigate this technique first for forecasts of Atlantic Multidecadal Variability (AMV, 146 see Methods). AMV is thought to be one of the most predictable aspects of decadal climate<sup>26</sup>, vet 147 the lagged ensemble mean does not capture the correct timing of the minimum in the late 1980s 148 (Figure 2c). NAO-matching captures the minimum and subsequent rapid warming in much bet-149 ter agreement with observations (Figure 2d) consistent with evidence that AMV is at least partly 150 forced by the NAO<sup>27–29</sup>. We find similar improvements for northern European rainfall: the lagged 151 ensemble mean is not significantly skilful and the observations lie outside the modelled uncer-152 tainties in the 1960s and 1980s (Figure 2e), whereas the NAO-matched ensemble is significantly 153 skilful (ACC=0.72, p<0.01) and captures the observed increase from the 1960s to late 1980s and 154 decrease thereafter. As expected, these improvements are not seen by simply adjusting the variance 155 of the ensemble mean (Supplementary Figure S1). 156

<sup>157</sup> Forecasts of extreme decades would be of particular value since they could enable action <sup>158</sup> to be taken in advance to avoid the most severe climate impacts<sup>30</sup>. We therefore investigate the <sup>159</sup> extreme positive NAO period between 1986 and 1997 (8-year means starting 1986 to 1990, Fig-<sup>160</sup> ure 2a). Consistent with the above results, the raw lagged ensemble mean shows virtually no signal <sup>161</sup> compared to observed variability (Figure 3 a, b, c compared to d, e, f). Adjusting its variance to be <sup>162</sup> equal to the observed variance (Figure 3 g, h, i) reveals that the forecasts do capture the positive <sup>163</sup> NAO (as expected from Figure 2b), but the expected impacts are underestimated, especially for <sup>164</sup> temperature and northern European precipitation. However, the NAO-matched forecast (Figure 3 <sup>165</sup> j, k, l) shows a clear improvement and captures the expected quadrupole pattern with warm, wet <sup>166</sup> (cold, dry) anomalies in northern Eurasia and south-east North America (northern Africa and parts <sup>167</sup> of southern Europe, and north-east North America), as well as low pressure across the Arctic. Sim-<sup>168</sup> ilar improvements from NAO-matching are found for trends and for skill measured over all of the <sup>169</sup> hindcasts (Supplementary Figures S3-S4).

We have shown that the winter NAO and related impacts on Europe and eastern North America are highly predictable on decadal timescales. AMV is usually believed to be a major source of decadal prediction skill<sup>26,31</sup>. However, we find that predictions of AMV can be improved by using the forecast NAO (Figure 2c,d), whereas predictions of the NAO are degraded by selecting the most skilful AMV ensemble members (Supplementary Figure S5). This suggests that the NAO is not solely driven by AMV. Hence other potential influences, including for example the tropics<sup>32–34</sup>, warrant further investigation.

Crucially we find that the NAO signal is underestimated by an order of magnitude in the model ensemble. This adds to an increasing body of evidence that the signal-to-noise ratio is too small in climate models, seen on seasonal<sup>20,35–37</sup>, interannual<sup>38</sup> and decadal<sup>19,39</sup> timescales. Consequently, the real world is more predictable than climate models suggest<sup>10,18</sup> and uncertainties diagnosed from raw model simulations are too large. The cause of this error is not yet known, though there are several hypotheses including weak teleconnections to the quasi-biennial oscillation<sup>40</sup>, lack of persistence in the NAO<sup>41,42</sup> and in weather regimes<sup>43</sup>, unresolved ocean at mosphere interactions<sup>44</sup> and weak transient eddy feedback<sup>45</sup>.

A key question is whether climate models also underestimate signals on timescales beyond a 185 decade. There is some evidence that the atmospheric circulation response to Arctic sea ice loss<sup>46</sup>, 186 and to external factors<sup>10</sup> including volcanic eruptions, solar variations and ozone changes, are too 187 weak in models. Models also appear to underestimate the magnitude of multi-decadal temperature 188 variability<sup>47,48</sup> especially for the north Atlantic<sup>49,50</sup>. Furthermore, model-simulated winter climate 189 change signals in the north Atlantic increase substantially as resolution increases<sup>51</sup>, consistent 190 with the suggestion that eddy feedbacks are inadequately resolved<sup>45</sup>. If this is robust, treating 191 current model simulations at face value is giving misleading conclusions about uncertainties and 192 irreducible internal variability. 193

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#### 197 Methods

Observations and models. Near surface temperature observations are computed as the average of HadCRUT4<sup>52</sup>, NASA-GISS<sup>53</sup> and NCDC<sup>54</sup>. Precipitation observations are taken from GPCC<sup>55</sup> and mean sea level pressure is taken from HadSLP $2^{56}$ .

We assess a large multi-model ensemble (169 members, Table 1) of decadal predictions from 201 14 modelling systems using hindcasts starting each year from 1960 to 2005 from the Coupled 202 Model Intercomparison Project (CMIP5) phases 5<sup>15</sup> and 6<sup>16</sup>. We found no significant difference 203 in NAO correlation skill between the CMIP5 and CMIP6 ensembles and focus on the combined 204 ensemble to obtain the most robust statistics. We create ensemble means by taking the equally-205 weighted average of all ensemble members and assess rolling 8-year boreal winter (December to 206 March) means defined by calendar years 2 to 9 from each start date. The forecasting systems 207 start between 1st of November and January each year, giving a lead time of at least a year before 208 the assessed forecast period to focus on decadal timescales and avoid predictability arising from 209 seasonal to annual variability. Both halves of the 8-year period contribute to skill (NAO ACC = 210 0.57 and 0.45, p=0.03, for forecast years 2 to 5 and 6 to 9 respectively). Both observations and 211 models were interpolated to a 5° longitude by 5° latitude grid before comparison. 212

Indices. The North Atlantic Oscillation (NAO) index is calculated as the difference in mean sea
level pressure between two small boxes located around the Azores (28-20°W, 36-40°N) and Iceland
(25-16°W, 63-70°N) with the average over the whole time series removed to create anomalies<sup>38</sup>.
Atlantic Multidecadal Variability (AMV) is calculated as the near-surface temperature in the North
Atlantic (80-0°W, 0-60°N) minus the global average (60°S-60°N)<sup>57</sup>. European rainfall is averaged
over the box 10°W-25°E, 55-70°N. All forecasts indices are based on the ensemble mean.

Forecast quality and uncertainty measures. Model biases and drifts are treated by computing anomalies relative to climatology for each model computed over all hindcasts, and comparing with observed anomalies computed over the same period. Although there are many ways to measure forecast quality, we focus on those that illustrate the underestimated model signals by using the following:

Pearson anomaly correlation coefficient ACC = 
$$\frac{\sum_{i=1}^{N} (f_i - \bar{f})(o_i - \bar{o})}{\sum_{i=1}^{N} (f_i - \bar{f})^2 \sum_{i=1}^{N} (o_i - \bar{o})^2}$$
(1)

Mean-squared-skill-score MSSS = 
$$1 - \frac{\sum_{i=1}^{N} (f_i - o_i)^2}{\sum_{i=1}^{N} (\bar{o} - o_i)^2}$$
 (2)

Ratio of predictable components RPC = 
$$\frac{\sigma_{sig}^o/\sigma_{tot}^o}{\sigma_{sig}^f/\sigma_{tot}^f} = \frac{ACC}{\sigma_{sig}^f/\sigma_{tot}^f}$$
 (3)

Ratio of predictable signals 
$$= \frac{\sigma_{sig}^o}{\sigma_{sig}^f} = RPC \frac{\sigma_{tot}^o}{\sigma_{tot}^f}$$
 (4)

where *N* is the number of hindcast start dates,  $f_i$  and  $o_i$  are the ensemble mean forecast and observations at each time, and the overbar represents the average over all times.  $\sigma_{sig}$  and  $\sigma_{tot}$  are the expected standard deviations of the predictable signal and total variability, with superscripts o and f for the observations and forecasts respectively. For the forecasts,  $\sigma_{sig}$  and  $\sigma_{tot}$  are computed from the ensemble mean and individual members respectively.

ACC measures the ability to predict the phase of variability, whereas MSSS measures the magnitude of errors relative to a climatological forecast. For a perfect forecasting system RPC should equal one. Note that RPC is not computed where the ACC is negative, and that the above formula likely gives a lower bound<sup>10, 18</sup>.

<sup>233</sup> Uncertainties in raw model forecasts are computed from the ensemble standard deviation

for each start date. Uncertainties in variance adjusted and NAO-matched forecasts are computed from the root-mean-square error between the ensemble mean and the observations as required for reliable forecasts<sup>58</sup>.

We note that it is theoretically possible for the multi-model RPC to be larger than for individ-237 ual models if time dependent model biases<sup>59</sup> or teleconnection errors reduce the model signal more 238 than the correlation with observations. Assessing this thoroughly would require large ensembles 239 of individual model hindcasts which are not available. However, assessing the largest individual 240 model ensemble available (NCAR CESM1.1 with 40 members per year, giving 160 lagged mem-241 bers, Table 1) does not support this hypothesis: the NCAR RPC of 6.2 is not significantly different 242 from the average RPC of multi-model ensembles of the same size (4.8 averaged over 1000 ran-243 dom samples, with 5-95% range 1.3 to 7.4). Furthermore, the statistics presented in this study are 244 appropriate for multi-model ensemble forecasts. 245

We further note that there is some evidence that the predictability of the NAO may vary on multi-decadal timescales<sup>60</sup>, though this is not robust across models<sup>61</sup>. Our results are statistically significant for the hindcast period available, but longer hindcasts that include more cycles of decadal variability would be beneficial for future studies.

Lagged ensemble. Consecutive 8-year means contain 7 identical years. Hence large interannual variations, as seen in 169-member ensemble mean NAO forecsts (Figure 2b), are not expected. They occur because the signal to noise ratio is too small in models and consecutive decadal predictions consist of independent model simulations that are dominated by different samples of the

noise. Ideally additional ensemble members would be used to reduce the noise further, but these 254 are not available. Instead we create a lagged ensemble by combining the required forecast with the 255 previous three i.e. the year 2-9 forecasts starting in 1963 are combined with the year 2-9 forecasts 256 starting in 1962, 1961 and 1960 giving a total of 676 members (169 members time 4 start dates). 257 The previous forecasts are sub-optimal because they do not cover exactly the same forecast period, 258 and rely on the persistence of running 8-year means. Hence there is a trade off between reducing 259 the noise with additional members and potentially degrading the skill by relying on persistence. 260 In the current generation of climate models the benefit in reducing the noise far outweighs the 261 degradation from using persistence. We present results for the combination of 4 lagged forecasts, 262 but find similar levels of skill for other combinations (NAO ACC = 0.71 and 0.78 for combining 3 263 and 5 lagged forecasts respectively). A similar technique relying on persistence of the predictor re-264 cently proved to strongly reduce the noise in decadal predictions of summer temperature extremes 265 over land<sup>62</sup>. 266

#### <sup>267</sup> NAO-matching. At any location that is influenced by the NAO we can write

$$O = O_{NAO} + O_{OTHER} + \epsilon^o \tag{5}$$

$$F^k = F^k_{NAO} + F^k_{OTHER} + \epsilon^k \tag{6}$$

$$\hat{F} = \hat{F}_{NAO} + \hat{F}_{OTHER} + \hat{\epsilon} \tag{7}$$

where O,  $F^k$  and  $\hat{F}$  are the observed, forecast ensemble member k and forecast ensemble mean values of a meteorological variable (e.g. temperature, rainfall, pressure). The subscript *NAO* refers to the portion that is related to the NAO, the subscript *OTHER* refers to the portion related to other predictable drivers (including greenhouse gases and sea surface temperatures unrelated to the NAO) and  $\epsilon$  is an unpredictable residual. Because the predictable NAO signal is too small in models, the mean of a very large ensemble is required for skilful NAO predictions (Figure 2b). However, the magnitude of the ensemble mean NAO is much too small (Figure 2a) and therefore  $\hat{F}_{NAO}$  will be severely underestimated.

One approach to overcoming model deficiencies uses regressions between model hindcasts and observations<sup>25,63–65</sup>, which effectively replaces the erroneous  $\hat{F}_{NAO}$  with the observed value  $O_{NAO}$ . Whilst this can give very good results, it relies on  $O_{NAO}$  estimated from the observations being robust and describing a causal relationship between the NAO and remote regions. This approach is less attractive on decadal than seasonal timescales because  $O_{NAO}$  is potentially more affected by sampling errors from the relatively small hindcast period.

An alternative approach<sup>24</sup> replaces the underestimated  $\hat{F}_{NAO}$  with more realistic  $F_{NAO}^k$  by 282 selecting from the full ensemble a smaller set of members that have the required magnitude of 283 the NAO. These members contain close-to the correct magnitude of the required NAO and its 284 teleconnections whilst retaining other influencies. Hence,  $\hat{F}_{NAO}$  for this selected ensemble will be 285 larger than that of the full ensemble, thereby increasing the signal. Because the selected ensemble 286 is smaller the remaining noise will not be reduced as much as in the full ensemble. However, 287 the selection process transfers variability from what would be considered as noise in a random 288 ensemble into  $\hat{F}_{NAO}$ , thereby reducing  $\hat{\epsilon}$  in the selected ensemble. Hence, in regions affected by 289 the NAO the increase in signal is likely to be larger than the reduced suppression of the remaining 290 noise, thereby increasing the signal to noise ratio and improving the skill. 29'

In the previous seasonal forecast study<sup>24</sup> the required NAO was obtained based on observed 292 relationships with potential drivers. However, on decadal timescales such relationships are not 293 well-established and are more likely to be affected by sampling errors. We therefore take the re-294 quired NAO to be the ensemble mean forecast NAO but adjusted to account for the underestimation 295 of the predictable signal. This is achieved by muliplying the ensemble mean NAO by the ratio of 296 predictable signals (equation 4). To avoid overfitting to observations we compute the ratio of pre-297 dictable signals for each hindcast start date separately using a cross-validation approach in which 298 the required hindcast and those on either side are omitted. Our conclusions are robust to omit-299 ting more hindcasts (we have tested up to 4 years either side) though skill may be underestimated 300 especially in these cases<sup>66,67</sup>. 30

#### <sup>302</sup> The overall procedure is as follows. For each start date i:

- 1. Compute the signal-adjusted (described above) NAO index of the ensemble mean  $NAO_i$
- <sup>304</sup> 2. Compute the NAO index for each ensemble member NAO<sup>k</sup><sub>i</sub>
- 305 3. For each ensemble member calculate the difference  $NAO_i^k N\hat{A}O_i$
- 4. Select the M (= 20) ensemble members with the smallest absolute differences

We take the mean of this subset of M members and present standardised forecast anomalies (Figure 3) or adjust its variance to be the same as observed (Figure 2). We note that this approach is applicable to forecasts as well as hindcasts. We present results for a subset of 20 members, but the results are similar for subsets ranging from 10 to 40 members. This method relies on models simulating realistic NAO teleconnections  $(F_{NAO}^k)$  and further improvements might be possible by using the best models in this respect, but this is beyond the scope of this study.

Significance. For a given set of validation cases, we test for values that are unlikely to be accounted for by uncertainties arising from a finite ensemble size (*E*) and a finite number of validation points (*N*). This is achieved using a non-parametric block bootstrap approach<sup>19,68,69</sup>, in which an additional 1000 hindcasts are created as follows:

1. Randomly sample with replacement N validation cases. In order to take autocorrelation into account this is done in blocks of 5 consecutive cases.

 $_{319}$  2. For each of these, randomly sample with replacement *E* ensemble members.

320 3. Compute the required statistic for the ensemble mean (e.g. correlation, MSSS, RPC).

4. Repeat from (1) 1000 times to create a probability distribution.

5. Obtain the significance level based on a 2-tailed test of the hypothesis that skill is zero, or RPC is one.

324 **Data Availability** The datasets analysed during the current study are available from the CMIP data archives.

Code Availability The code used during the current study is available from the corresponding author on reasonable request.

Acknowledgements DMS, AAS, NJD, LH and RE were supported by the Met Office Hadley Centre Climate Programme funded by BEIS and Defra and by the European Commission Horizon 2020 EUCP

project (GA 776613). FJDR, LPC, SW and RB also acknowledge the support from the EUCP project 329 (GA 776613) and from the Ministerio de Economía y Competitividad (MINECO) as part of the CLINSA 330 project (Grant No. CGL2017-85791-R). SW received funding from the innovation programme under the 331 Marie Sklodowska-Curie grant agreement H2020-MSCA-COFUND-2016-754433 and PO from the Ramon 332 y Cajal senior tenure programme of MINECO. The EC-Earth simulations were performed on Marenostrum 333 4 (hosted by the Barcelona Supercomputing Center, Spain) using Auto-Submit through computing hours 334 provided by PRACE. WAM, HP, KM and KP were supported by the German Federal Ministry for Education 335 and Research (BMBF) project MiKlip (grant 01LP1519A). NK, IB, FC and YW were supported by the 336 Norwegian Research Council projects SFE (grant 270733) the Nordic Center of excellent ARCPATH (grant 337 76654) and the Trond Mohn Foundation, under the project number : BFS2018TMT01 and received grants 338 for computer time from the Norwegian Program for supercomputing (NOTUR2, NN9039K) and storage 339 grants (NORSTORE, NS9039K). JM, LFB and DS are supported by Blue-Action (European Union Horizon 340 2020 research and innovation program, Grant Number: 727852) and EUCP (European Union Horizon 2020 341 research and innovation programme under grant agreement no 776613) projects. The National Center for 342 Atmospheric Research (NCAR) is a major facility sponsored by the US National Science Foundation (NSF) 343 under Cooperative Agreement No. 1852977. NCAR contribution was partially supported by the National 344 Oceanic and Atmospheric Administration (NOAA) Climate Program Office under Climate Variability and 345 Predictability Program Grant NA13OAR4310138 and by the US NSF Collaborative Research EaSM2 Grant 346 OCE-1243015. 347

348 **Competing Interests** The authors declare that there are no competing interests.

Correspondence Correspondence and requests for materials should be addressed to D.M.S.
 (email: doug.smith@metoffice.gov.uk).

351	Author contributions	D.M.S. led the analysis and writing with comments from all authors.	R.E. pro-
352	cessed the CMIP5 data.	A.A.S. suggested NAO-matching. All authors except A.A.S., P.A., A.B.	., PA.M.,
353	D.N., J.R. and P.R. cont	ributed to creating the decadal prediction data.	

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Forecast Centre	Model	Atmosphere	Ocean resolution <sup>2</sup>	Ensemble	CMIP	
		resolution <sup>1</sup>		size	version	
Barcelona Supercomputing Center,	EC-Earth3 <sup>70,71</sup>	0.7x0.7x91x0.01	1x1x0.3x75	10	CMIP6	
Spain						
Bjerknes Center for Climate Research,	NorCPM1 <sup>72,73</sup>	1.9x2.5x26x3	0.7x1.125x0.25x53	20	CMIP6	
Norway						
Canadian Centre for Climate Modelling	CanCM4 <sup>74</sup>	2.8x2.8x35x1	0.94x1.41x40	10	CMIP5	
and Analysis, Environment and Climate						
Change Canada						
	CanESM5 <sup>75, 76</sup>	2.8x2.8x49x1	1x1x0.3x45	10	CMIP6	
Geophysical Fluid Dynamics Laboratory,	CM2.1 <sup>77</sup>	2x2.5x24x3	1x1x0.3x50	10	CMIP5	
USA						
IPSL-EPOC, France	IPSL-CM6A-LR	1.25x2.5x79x0.005	1x1x0.3x75	10	CMIP6	
Met Office Hadley Centre, UK	HadCM3 <sup>67</sup>	2.5x3.75x19x4.5	1.25x1.25x20	20	CMIP5	
	HadGEM3 <sup>78</sup>	0.55x0.83x85x0.005	0.25x0.25x75	10	CMIP6	
Max Planck Institute for Meteorology,	MPI-ESM1.0-LR <sup>79</sup>	1.9x1.9x47x0.01	1.5x1.5x40	3	CMIP5	
Germany						
	MPI-ESM1.2-	0.9x0.9x95x0.01	0.4x0.4x40	10	CMIP6	
	HR <sup>80</sup>					
National Center for Atmospheric Re-	CESM1.1 <sup>39</sup>	0.9x1.25x30x2.26	1x1.125x0.27x60	40	CMIP6	
search, USA						
University of Tokyo, National Institute	MIROC5 <sup>81,82</sup>	1.4x1.4x40x3	1.4x1.4x0.5x49	6	CMIP5	
for Environmental Studies, and Japan						
Agency for Marine-Earth Science and						
Technology, Japan						
	MIROC6	1.4x1.4x81x0.004	1x1x0.5x62	10	CMIP6	

#### Table 1: Forecast systems and ensemble sizes.

<sup>1</sup> Atmosphere resolution (degrees latitude)x(degrees longitude)x(number of vertical levels)x(lid height, hPa)

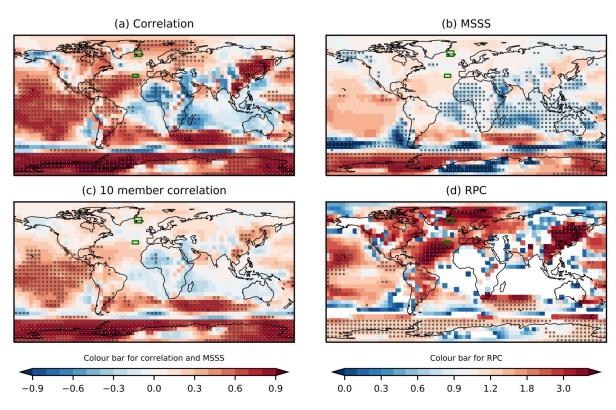
<sup>2</sup> Ocean resolution (degrees latitude)x(degrees longitude)x(optional degrees latitude at Equator)x(number of vertical levels)

Decadal prediction skill for boreal winter (December to March) mean sea Figure 1 538 level pressure. Skill for year 2-9 multi-model ensemble mean forecasts measured by (a) 539 anomaly correlation, (b) mean squared skill score (MSSS), (c) average anomaly correla-540 tion for a 10-member ensemble mean (computed over 1000 random samples). (d) The 541 ratio of predictable components (RPC). RPC is not calculated where the correlation is 542 negative. Stippling shows where correlations and MSSS, or RPC, are significantly dif-543 ferent to zero, or greater than one, respectively (95% confidence interval, see Methods). 544 Green boxes show the regions used to calculate the NAO. 545

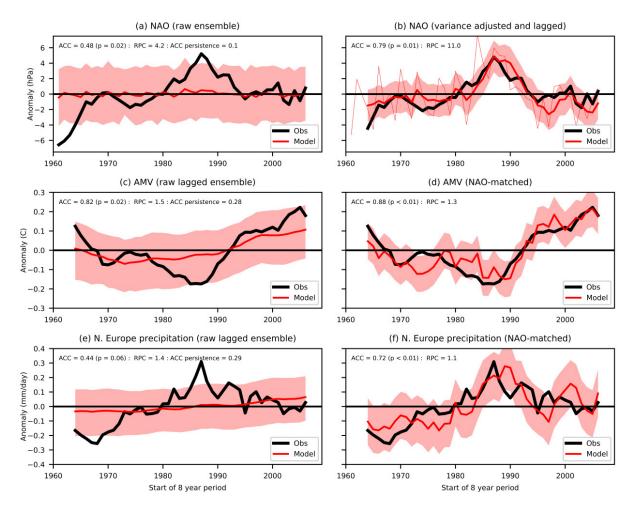
Figure 2 Underestimated signals. (a) Time series of observed (black curve) and model 546 forecast (years 2-9, red curve showing ensemble mean of 169 members and red shading 547 showing the 5-95% confidence interval diagnosed from the individual members) 8-year 548 running mean December to March NAO index. (b) As (a) but for ensemble mean forecast 549 rescaled to have the same variance as the observations (thin red curve), and addition-550 ally smoothed by taking the lagged average of the latest four forecasts at each start date 551 (thick red curve, 676 members, see Methods). Forecast uncertainty (red shading, 5-95%) 552 confidence interval) is obtained from the forecast ensemble mean error variance (see 553 Methods). (c) As (a) but for AMV and lagged ensemble. (d) As (c) but for NAO-matched 554 forecast (see Methods). (e, f) As (c, d) but for northern European rainfall. Values of 555 anomaly correlation (ACC) of the forecast ensemble mean and of persisting the latest 556

observed 8-year mean available before each start date, and the ratio of predictable com ponents (RPC), are indicated. Indices are defined in Methods. Time-series are anomalies
 relative to the average of all year 2-9 hindcasts.

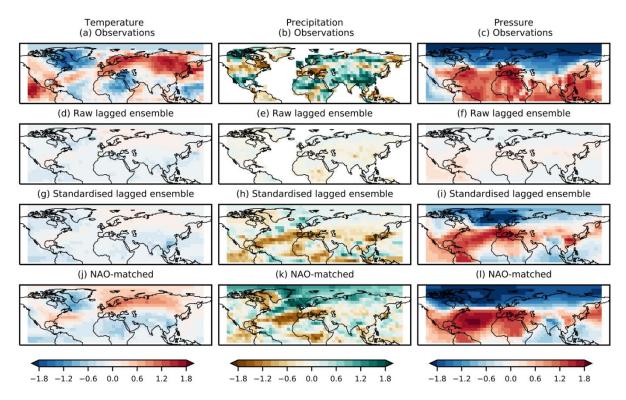
**Figure 3** Decadal predictions of the extreme NAO period (1986 to 1997). Observed 560 anomalies of (a) temperature, (b) precipitation and (c) mean sea level pressure. (d, e, 561 f) As (a, b, c) but for raw lagged ensemble mean forecasts. (g, h, i) As (d, e, f) but 562 standardised by the ensemble mean standard deviation. (j, k, l) As (d, e, f) but for NAO-563 matched forecasts. Averages are taken for boreal winter (December to March) for all year 564 2-9 forecasts verifying in the period 1986 to 1997 (i.e. start dates 1985 to 1989 inclusive), 565 and converted to anomalies by removing the average over all hindcasts (i.e. start dates 566 1960 to 2005 inclusive). Units are standard deviations. The raw lagged ensemble (d, e, 567 f) is divided by the observed standard deviation to show the signal relative to observed 568 variability. 569



**Figure 1: Decadal prediction skill for boreal winter (December to March) mean sea level pressure.** Skill for year 2-9 multi-model ensemble mean forecasts measured by (a) anomaly correlation, (b) mean squared skill score (MSSS), (c) average anomaly correlation for a 10-member ensemble mean (computed over 1000 random samples). (d) The ratio of predictable components (RPC). RPC is not calculated where the correlation is negative. Stippling shows where correlations and MSSS, or RPC, are significantly different to zero, or greater than one, respectively (95% confidence interval, see Methods). Green boxes show the regions used to calculate the NAO.



**Figure 2: Underestimated signals.** (a) Time series of observed (black curve) and model forecast (years 2-9, red curve showing ensemble mean of 169 members and red shading showing the 5-95% confidence interval diagnosed from the individual members) 8-year running mean December to March NAO index. (b) As (a) but for ensemble mean forecast rescaled to have the same variance as the observations (thin red curve), and additionally smoothed by taking the lagged average of the latest four forecasts at each start date (thick red curve, 676 members, see Methods). Forecast uncertainty (red shading, 5-95% confidence interval) is obtained from the forecast ensemble mean error variance (see Methods). (c) As (a) but for AMV and lagged ensemble. (d) As (c) but for NAO-matched forecast (see Methods). (e, f) As (c, d) but for northern European rainfall. Values of anomaly correlation (ACC) of the forecast ensemble mean and of persisting the latest observed 8-year mean available before each start date, and the ratio of predictable components (RPC), are indicated. Indices are defined in Methods. Time-series are anomalies relative to the average of all year 2-9 hindcasts.



**Figure 3: Decadal predictions of the extreme NAO period (1986 to 1997).** Observed anomalies of (a) temperature, (b) precipitation and (c) mean sea level pressure. (d, e, f) As (a, b, c) but for raw lagged ensemble mean forecasts. (g, h, i) As (d, e, f) but standardised by the ensemble mean standard deviation. (j, k, l) As (d, e, f) but for NAO-matched forecasts. Averages are taken for boreal winter (December to March) for all year 2-9 forecasts verifying in the period 1986 to 1997 (i.e. start dates 1985 to 1989 inclusive), and converted to anomalies by removing the average over all hindcasts (i.e. start dates 1960 to 2005 inclusive). Units are standard deviations. The raw lagged ensemble (d, e, f) is divided by the observed standard deviation to show the signal relative to observed variability.