

Empirical likelihood tests for nonparametric detection of differential expression from RNA-seq data

Article

Supplemental Material

Thorne, T. (2015) Empirical likelihood tests for nonparametric detection of differential expression from RNA-seq data. Statistical Applications in Genetics and Molecular Biology, 14 (6). pp. 575-583. ISSN 2194-6302 doi: https://doi.org/10.1515/sagmb-2015-0095 Available at https://centaur.reading.ac.uk/73955/

It is advisable to refer to the publisher's version if you intend to cite from the work. See <u>Guidance on citing</u>. Published version at: http://dx.doi.org/10.1515/sagmb-2015-0095 To link to this article DOI: http://dx.doi.org/10.1515/sagmb-2015-0095

Publisher: De Gruyter

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Supplementary material

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November 6, 2015

1 Empirical likelihood

For n i.i.d one dimensional observations x_1, \ldots, x_n the empirical likelihood (Owen, 1988) can be defined as

$$f(x) = \prod_{i=1}^{n} p_i,\tag{1}$$

where we assign each observation a weight p_i , and constrain these such that $\sum_{i=1}^{n} p_i = 1, \forall i, 0 \le p_i \le 1$. Focussing on the empirical likelihood for the mean μ of our observations x_i , we simply require that

$$\sum_{i=1}^{n} p_i x_i = \mu. \tag{2}$$

Then we have three constraints, and aim to find the the p_i that maximise the empirical likelihood f(x) under these constraints. Fortunately by using Lagrange multipliers we can find the optimal p_i by solving a one dimensional root finding problem. Defining

$$G = \sum_{i=1}^{n} \log(np_i) - n\lambda \sum_{i=1}^{n} p_i(x_i - \mu) + \gamma \left(\sum_{i=1}^{n} p_i - 1\right),$$
(3)

and taking the partial derivative with respect to p_i , applying the method of Lagrange multipliers (Owen, 2001) we have

$$\frac{\partial G}{\partial p_i} = \frac{1}{p_i} - n\lambda(x_i - \mu) + \gamma = 0, \tag{4}$$

and we can solve for γ by considering

$$\sum_{i=1}^{n} p_i \frac{\partial G}{\partial p_i} = 0 \tag{5}$$

$$\sum_{i=1}^{n} (1 - n\lambda p_i(x_i - \mu) + p_i\gamma) = 0$$
 (6)

$$n + \gamma = 0, \tag{7}$$

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since we know $\sum_{i=1}^{n} p_i(x_i - \mu) = 0$. Then substituting $\gamma = -n$ into equation 4 we have

$$\frac{1}{p_i} - n\lambda(x_i - \mu) - n = 0 \tag{8}$$

$$p_i = \frac{1}{n\lambda(x_i - \mu) + n},\tag{9}$$

and so p_i depends only on solving equation 4 for λ . We know that

$$\sum_{i=1}^{n} p_i(x_i - \mu) = 0 \tag{10}$$

$$\sum_{i=1}^{n} \frac{(x_i - \mu)}{n\lambda(x_i - \mu) + n} = 0, \qquad (11)$$

and so we can solve for λ for a given value of μ using a univariate root finding algorithm. Then using equation 9 we can find the p_i and calculate the empirical likelihood in equation 1.

1.1 Euclidean likelihood

The Euclidean likelihood (Baggerly, 1998) defines the log likelihood as

$$\log f(x|\mu) = -\frac{1}{2} \sum_{i=1}^{n} (np_i - 1)^2,$$
(12)

with the constraints $\sum_{i=1}^{n} p_i = 1$ and $\sum_{i=1}^{n} p_i x_i - \mu = 0$. Again we apply the method of Lagrange multipliers (Owen, 2001)

$$G = -\frac{1}{2} \sum_{i=1}^{n} (np_i - 1)^2 - n\lambda \sum_{i=1}^{n} p_i(x_i - \mu) + \gamma \left(\sum_{i=1}^{n} p_i - 1\right), \quad (13)$$

and setting the partial derivative of G with respect to p_i to zero we have

$$\frac{\partial G}{\partial p_i} = n(1 - np_i) - n\lambda(x_i - \mu) + \gamma = 0$$
(14)

$$\frac{1}{n}\sum_{i=1}^{n} \left(n(1-np_i) - n\lambda(x_i - \mu) + \gamma\right) = 0$$
(15)

$$-n\lambda(\bar{x}-\mu)+\gamma = 0.$$
(16)

Substituting $\gamma = n\lambda(\bar{x} - \mu)$ back into equation 14

$$n(1 - np_i) - n\lambda(x_i - \mu) + n\lambda(\bar{x} - \mu) = 0$$
(17)

$$p_i = \frac{1}{n} (1 - \lambda (x_i - \bar{x})).$$
 (18)

Given that $\sum_{i=1}^{n} p_i(x_i - \mu) = 0$, we can substitute equation 18 to give

$$\sum_{i=1}^{n} \frac{(x_i - \mu)}{n} \left(1 - \lambda (x_i - \bar{x}) \right) = 0$$
(19)

$$\bar{x} - \mu - \sum_{i=1}^{n} \frac{\lambda}{n} (x_i - \mu) (x_i - \bar{x}) = 0$$
(20)

$$\bar{x} - \mu - \sum_{i=1}^{n} \frac{\lambda}{n} (x_i - \bar{x}) (x_i - \bar{x}) = 0$$
(21)

$$\bar{x} - \mu - \lambda s = 0, \qquad (22)$$

where s is defined as $s = \frac{1}{n} \sum_{i=1}^{n} (x_i - \bar{x})(x_i - \bar{x})$. Substituting λ into equation 18 we have

$$p_i = \frac{1}{n} \left(1 - \frac{1}{s} (\bar{x} - \mu) (x_i - \bar{x}) \right), \tag{23}$$

and substituting p_i into equation 12 we arrive at

$$\log f(x|\mu) = -\sum_{i=1}^{n} \left(\frac{1}{s}(\bar{x}-\mu)(x_i-\bar{x})\right)^2$$
(24)

$$= -\frac{1}{s^2}(\bar{x}-\mu)^2 \left(\sum_{i=1}^n (x_i-\bar{x})^2\right)$$
(25)

$$= -\frac{1}{s}n(\bar{x}-\mu)^2,$$
 (26)

References

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